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**A Coincident Economic Indicator
of Economic Activity in Greece**

Ekaterini Tsouma

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Ekaterini Tsouma

Research Fellow

Centre for Planning and Economic Research (KEPE), Athens, Greece

e-mail for correspondence: etsouma@kepe.gr

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Οικονομικός Δείκτης Συγκυρίας για την Οικονομική Δραστηριότητα στην Ελλάδα

Αικατερίνη Τσούμα

ΠΕΡΙΛΗΨΗ

Η κατασκευή και ανάλυση σύνθετων οικονομικών δεικτών συγκυρίας για την έγκαιρη αναγνώριση της τρέχουσας οικονομικής συγκυρίας αποτελεί πλέον παράδοση στις ΗΠΑ και άλλες ανεπτυγμένες οικονομίες. Με στόχο την απόκτηση ενός αξιόπιστου και έγκαιρα διαθέσιμου συνολικού μεγέθους σε μηνιαία συχνότητα για την αξιολόγηση της τρέχουσας οικονομικής δραστηριότητας, στην παρούσα εργασία εισάγεται ένας σύνθετος οικονομικός δείκτης συγκυρίας για την Ελληνική οικονομία. Ο προτεινόμενος δείκτης κατασκευάζεται στην βάση των δύο προσδιοριστικών χαρακτηριστικών των οικονομικών διακυμάνσεων: Του ότι οι μεμονωμένες οικονομικές μεταβλητές δείχνουν να συμβαδίζουν και ότι παρατηρείται ασύμμετρη συμπεριφορά μεταξύ οικονομικών υφέσεων και επεκτάσεων. Ο σύνθετος οικονομικός δείκτης συγκυρίας για την Ελληνική οικονομία κατασκευάζεται μέσω της εφαρμογής της συνθετικής μεθοδολογίας του υποδείγματος δυναμικού παράγοντα και εναλλαγής καθεστώτος. Η συγκεκριμένη μεθοδολογία επιλέγεται διότι καθιστά δυνατή την απόκτηση μέσα από το υπόδειγμα όχι μόνο του σύνθετου δείκτη συγκυρίας ως μεμονωμένης μεταβλητής που αντανακλά την 'κατάσταση της οικονομίας', αλλά και των πιθανοτήτων ύφεσης και επέκτασης. Το επιλεγμένο υπόδειγμα εκτιμάται χρησιμοποιώντας μεμονωμένες οικονομικές μεταβλητές σε μηνιαία συχνότητα και η περίοδος κάλυψης είναι από τον Ιανουάριο 1970 έως τον Δεκέμβριο 2007. Τα εμπειρικά αποτελέσματα στηρίζουν την εφαρμογή του συνθετικού υποδείγματος επιβεβαιώνοντας την ασυμμετρία μεταξύ υφέσεων και επεκτάσεων και παρέχουν χρήσιμη πληροφόρηση για τις κυκλικές διακυμάνσεις στην Ελληνική οικονομική δραστηριότητα. Ο σύνθετος οικονομικός δείκτης συγκυρίας και οι πιθανότητες ύφεσης που εξάγονται από την εκτίμηση του συνθετικού υποδείγματος φαίνεται να απεικονίζουν ικανοποιητικά τους οικονομικούς κύκλους για την Ελλάδα κατά την υπό εξέταση χρονική περίοδο.

ABSTRACT

In the USA and other advanced economies, it has long been a tradition to develop and analyze composite coincident economic indicators to provide a broad representation of the underlying economic conditions. In order to develop a similar indicator of the direction of economic activity in Greece on a monthly basis, in this paper we introduce a composite coincident indicator of Greek economic activity. The proposed indicator is developed on the basis of the two determining features of economic fluctuations: co-movement among individual economic variables and asymmetric behavior in recessions and expansions. The indicator is constructed by applying a dynamic factor model with regime switching. The selected methodology allows us to obtain both the composite coincident indicator as a single unobserved variable, and the implied probabilities for the regimes of expansion and recession. We use monthly data covering the January 1970-December 2007 period. The evidence supports the application of the Markov switching framework. The empirical results indicate that the applied model with regime switching leads to a satisfactory representation of the sample data. The estimated coincident indicator appears to be adequate for an evaluation of Greek economic activity, since stylized facts of the Greek business cycle are well captured by the model.

1. Introduction

Business and economic policy decisions require a timely understanding of the current status of aggregate economic activity. However, in most cases no broad measure indicating the direction of economic activity is available on a monthly basis. To respond to this need, it has long become a tradition in the USA and other advanced economies to develop and analyze composite coincident economic indicators. Such research was initiated by Burns and Mitchell (1947) for the USA. It relies on the concept that the combination of the information incorporated in individual economic series into a composite indicator can provide a broad representation of the underlying economic conditions.

In Greece, where there is also no available measure of aggregate economic activity on a monthly basis, the existence of a timely available and reliable composite coincident economic indicator is essential for economic policy and business decision making. Hence, in this paper we propose a composite coincident indicator of economic activity in Greece on the basis of the two determining features of economic fluctuations: co-movement among individual economic variables and asymmetric behavior in recessions and expansions. We apply the dynamic factor model methodology with regime switching according to Diebold and Rudebusch (1996), Chauvet (1998) and Kim and Nelson (1999). Within this framework, we obtain both the composite coincident indicator as a single unobserved variable, “the state of the economy”, and the implied probabilities of expansion and recession. To our knowledge, this is the first attempt to construct a composite coincident indicator of Greek economic activity using the selected methodology. We think such an extension

of the existing empirical literature, which is indeed very limited, is important for two basic reasons: first, the combination of the individual variables and the derivation of the indicator rely on a formal econometric approach and, second, the methodology applied provides us with signals of changes in aggregate economic activity. The results support the application of the specific methodology and stylized facts of the Greek business cycle are well captured by the model.

The rest of the paper is organized as follows. Section 2 provides a review of the relevant literature. Section 3 outlines the underlying empirical methodology and presents the applied model specification. Section 4 discusses the data and Section 5 presents the empirical results. The last section concludes.

2. A Review of the Literature

Since the introduction of cyclical indicators for monitoring economic conditions in the USA by Burns and Mitchell in the 1920s and 1930s, coincident indicators¹ have been widely used to indicate the direction of economic activity worldwide. The pioneering work of Burns and Mitchell (1947) has led, on the one hand, to serious criticism such as the Koopmans's (1947) "measurement without theory" critique; on the other hand, it has spurred the production of a wide literature dealing with the development of sophisticated methods for the combination of individual series into composite indices and the construction of coincident indicators of economic activity.

¹ According to whether the fluctuations in the monitored series persistently led, coincided or lagged turning points in US BC, Burns and Mitchell classified economic variables as leading, coincident and lagging, respectively.

Composite indicators for the US economy were first published in November 1968.² The construction of the composite coincident index of economic activity was based on a weighting scheme developed and applied by Moore and Shiskin (1967)³. To date, the official coincident index for US economic activity released by the Conference Board is calculated as a weighted average of four well-established, broad-based, timely series, namely, industrial production, personal income less transfer payments, employees on non-agricultural payrolls and manufacturing and trade sales. The underlying weighting procedure has been extensively used until the late 1980s. It still forms the basis for numerous applications for the construction of coincident indices of economic activity,⁴ such as the work of Layton and Moore (1989) for the US service sector, Phillips, Vargas, and Zarnowitz (1996) for the Mexican economy, Altissimo, Marchetti, and Oneto (2000) for the Italian economy and Lamy and Sabourin (2001) for Canada. Despite the popularity of the weighting approach and the satisfactory results it has provided, there was still the argument that the associated weights were chosen subjectively and the analysis did not rely on any formal econometric framework.

² They were published by the Bureau of the Census. of the US Department of Commerce in the Business Conditions Digest. In 1968, the original monthly report *Business Cycle Developments* was renamed *Business Conditions Digest*. The publication of the first was initiated in 1961 by the US Government in cooperation with the NBER (see Business Cycle Indicators Handbook, Conference Board, 2000). In 1972, the indicators were shifted to the Bureau of Economic Analysis and in December 1995 the release of the indicators was taken over by the Conference Board.

³ According to that scheme, variables were scored based on specific criteria. These criteria were economic significance, statistical adequacy, cyclical timing and business cycle conformity. Composite indicators for the US economy were first published in November 1968 by the Bureau of the Census of the US Department of Commerce in the Business Conditions Digest. In December 1995 the release of the indicators was taken over by the Conference Board.

⁴ For a comprehensive survey on non-model based and model based approaches for the construction of composite coincident indicators, see the reference to the choice and construction of the target variable that the leading indicators are supposed to lead in Marcellino (2005). The author concludes that the estimation of the current economic condition is rather robust to the choice of methodology. In a paper reviewing the main statistical techniques for constructing coincident indicators, Carriero and Marcellino (2007) come to similar conclusions on the comparability of the derived coincident indicators.

To face such criticism, Stock and Watson (1988a, 1989) developed an econometric approach to derive, among other indices, a new composite coincident indicator for US economic activity. As a result, the work of Stock and Watson (hereafter S&W) significantly enriched the analysis of cyclical composite indicators of economic activity. Within the framework of an explicit probability model, the dynamic factor model, the composite coincident indicator is derived as a single unobserved variable, “the state of the economy”. The constructed index, which is quantitatively similar to the DOC index, is based on four series: industrial production, real personal income less transfer payments, real manufacturing and trade sales, and employee-hours in non-agricultural establishments. The S&W dynamic factor model became very popular and caused a surge in related work, mostly in/for the USA and several advanced economies.⁵ The S&W approach has been also applied by researchers to construct coincident indices on the regional or state level, but also for specific sector economic activity. Various modifications and extensions to the basic S&W model have been also developed.⁶

An important extension to the basic S&W model is introduced in Mariano and Murasawa (2003), who develop a factor model of time series with mixed frequencies. In an attempt to improve the basic factor model, they include quarterly real GDP in order to take into account the important information contained therein. On the basis of this earlier work, Mariano and Murasawa (2004) derive a two-factor model for the

⁵ Similar applications can be found in Kim and Nelson (1999) for the US, Gaudreault, Lamy and Liu (2003) for the Canadian economy, Fukuda and Onodera (2001) for Japan, Chen (2007), in research on the synchronization of the Eurozone Business Cycles, for the Eurozone aggregate, individual Eurozone countries (such as Germany, France, Austria, Belgium, Italy, the Netherlands, Spain, Finland, Portugal) and to the USA, UK and Canada, and Hall and Zonzilos (2003) for the Greek economy.

⁶ Clayton-Matthews and Stock (1998) use the S&W methodology to estimate a coincident index for the Massachusetts economy, Megna and Xu (2003) for the New York State economy, Phillips (2004) for Texas economic activity, Crone and Clayton-Matthews (2004) for all 50 states and Cañas, Gilmer, and Phillips (2003) for Houston. Lahiri, Yao, and Young (2003) and Lahiri and Yao (2004) apply the S&W dynamic factor model to the transportation sector of the US economy using sectoral data. Using a modification to the basic S&W dynamic factor model, Nieto and Melo (2001) construct a coincident indicator for the Colombian economy.

observable mixed-frequency series. Crone and Clayton-Matthews (2004) also consider an extension with mixed frequency series in the construction of coincident indices for the 50 states. In an application of dynamic factor analysis to construct coincident indices for the USA and the Euro area, Proietti and Moauro (2006) take into account the non-linear temporal aggregation constraint when using mixed-frequency data. Kholodilin (2002) extends the basic S&W one-factor model to a two-factor model of the common coincident and leading economic indicators, where the two factors are estimated simultaneously.

Limitations faced regarding large data sets due to computational problems and the need to identify and select the coincident variables to be included in the model prior to estimation (basically by heuristic criteria) are the most cited drawbacks of the basic S&W methodology. To cope with these problems, Forni, Hallin, Lippi, and Reichlin (FHLR 2000a), FHLR (2000b, 2001) and S&W (2002a, 2002b) developed alternative procedures for the analysis of dynamic factor models. FHLR (2000a) introduced the generalized dynamic factor model which relates/reconciles dynamic factor analysis with dynamic principal component analysis. The model underlying the method for the construction of the EuroCOIN coincident indicator compiled for the Euro area belongs also to the class of FHLR models.⁷ In a recent work on the Belgian economy using a large data set, Van Nieuwenhuyze (2006) also applies the specific version of the FHLR generalized dynamic factor model and derives (among others) a common component driven by the Belgian business cycle.

⁷ The project for the construction of a coincident indicator for the Euro area, as described in Altissimo et al. (2001), has led to regular monthly release of the EuroCOIN coincident indicator since January 2002. The innovation introduced in the FHLR (2005) model also initiated refinements in the procedure applied for the production of the real time EuroCOIN indicator (Altissimo et al. (2006)), so that no future information be/is required.

The basic S&W dynamic factor model and all the related extensions and modifications which were developed thereafter incorporated one key attribute outlined in Burns and Mitchell's definition of business cycles: co-movement among individual economic variables. However, the same did not hold for the second prominent feature, i.e. the separate treatment of expansions and contractions. This key element was considered by Hamilton (1989) in a non-linear model for real GNP with discrete regime switching between periods of expansion and contraction. In an effort to combine the two main ideas used by scholars of the cycle, Diebold and Rudebusch (1996) argue that the dynamic factor and regime switching literatures should not be considered in isolation. They propose an empirical synthesis and describe a framework for the analysis of business cycle data incorporating both factor structure and regime switching,⁸ even though they do not fully estimate the synthesized model. The task of estimating such a model was tackled by Kim and Yoo (1995) and Chauvet (1998) who used filtering and smoothing methods developed by Kim (1994) for the maximum likelihood estimation. Kim and Nelson (1998a, 1998b) also apply dynamic factor models with regime switching relying on different estimation procedures.

Similarly to the basic S&W factor model, the synthesized dynamic factor model with regime switching caused a surge in similar applications, as well as extensions and modifications.⁹ Kholodilin (2003) considers a similar model (among others) for the Japanese business cycle and in extending earlier work he proposes and estimates (Kholodilin (2001)) a dynamic factor model with regime switching with mixed frequencies for the US economy and, furthermore, introduces and estimates a two-factor model which takes into account coincident as well as leading variables

⁸ They also offer an extensive survey of related literature and provide some links between macroeconomic theory and factor structure and regime switching.

⁹ Such applications refer also to the construction and estimation of composite leading indicators. See for example Bandholz and Funke (2003) for Germany and Bandholz (2005) for Hungary and Poland.

(Kholodilin (2002)). Similar models are applied to the German and to large European economies in Kholodilin (2005) and Kholodilin (2006), respectively. In an application to the UK economy, Mills and Wang (2003) propose and estimate a common factor regime switching. In his work mentioned above investigating the synchronization of the Eurozone Business Cycles, Chen (2007) further applies the common factor model with regime switching. Finally, in applications to the US transportation sector, Lahiri, Yao, and Young (2003) and Lahiri and Yao (2004) apply dynamic factor models with regime switching.

3. Data

We use five monthly coincident economic variables to estimate the above described model.¹⁰ The selected series, as described in Table 1, are chosen to provide a broad coverage of the underlying conditions in the Greek economy. The selection of variables is further dictated by data availability considerations and is also based on statistical criteria.¹¹ The data sample covers the time period from January 1970 to December 2007 and includes the following series: (a) industrial production, (b) retail sales, (c) imports, (d) tourism (arrivals of foreigners) (e) and passenger cars (licenses issued). We believe the series under consideration to be satisfactorily diversified in

¹⁰ Nominal variables are deflated and, wherever necessary, series are seasonally adjusted using the US Census Bureau X12 procedure. See also Table 1.

¹¹ Although data on employment and/or unemployment and personal income belong to the basic components used to construct coincident economic indicators for the US and other advanced economies, we do not include such data due to data availability restrictions. More generally, it is noted that the selection of the data series to be used for model estimation is highly constrained by the restricted availability of historical monthly data. An alternative would be the examination of a considerably shorter time period, possibly enlarging the set of available monthly series. However, it is believed that the validity and robustness of the indicator should be in a first stage established by showing its consistence with historical stylized facts of the Greek business cycle.

their coverage of the Greek economic sectors so that the index to be constructed can be expected to capture the cyclical movements in aggregate economic activity.¹²

To evaluate the usefulness of the proposed variables as potential coincident series, we check for their contemporaneous correlation with GDP.¹³ For all series but imports the correlations at lag zero are significantly high. Furthermore, the tourism series appears to exhibit a slightly lagging structure. Moreover, we do not find high correlations between industrial production and the other four series and, as a result we do not expect the estimated indicator to be dominated by movements in industrial production. We further conduct ADF unit root tests to check for stationarity and cointegration tests to check for cointegration. The test results are presented in Table 2. In all cases we cannot reject the null hypothesis of the existence of a unit root for the series in levels but the test results suggest rejection of the null when first differences are considered. Moreover, according to the conducted Johansen cointegration test we cannot reject the hypothesis of no cointegration. As a result, the model is estimated in first differences of the logarithms of the selected variables.

4. Empirical methodology and model specification

In this paper we apply the dynamic factor model with regime switching to develop a coincident index of economic activity for the Greek economy. To the best of our knowledge, this is the first attempt to construct such a coincident indicator for Greece on the basis of the specific methodology. The selection of this methodology is

¹² Our data sources are the General Secretariat of National Statistical Service of Greece and Eurostat.

¹³ Note, that we use real GDP quarterly data which are converted into monthly data (via a low-to-high frequency conversion procedure applied in EViews) as a benchmark series for the GDP series. It should be further noted that, against the background of the significant 2007 upward revision of GDP the General Secretariat of National Statistical Service of Greece provides adjusted GDP data only back to 2000. In order to construct a continuous GDP time series, we adjust the data backwards on the basis of the q-o-q percentage changes of the non-revised earlier quarterly data provided by Eurostat back to 1970.

based on considerations relating to the ability of the model to encompass the two defining features of business cycles and to provide us with implied probabilities of recessions and expansions which can be used to date turning points.

In the following model which synthesizes the S&W (1989) dynamic factor model with Hamilton's (1989) regime-switching model, the coincident indicator is extracted as the common component out of a group of exogenous economic variables signaling the current state of the economy. At the same time, the cyclical movement of the indicator is translated into probabilities of being in a recession or expansion. Defining $\Delta y_{it} = \Delta Y_{it} - \Delta \bar{Y}_i$ and $\Delta c_t = \Delta C_t - \delta$, the applied model is expressed in deviations from means as follows:¹⁴

$$\Delta y_{it} = \gamma_i(L)\Delta c_t + e_{it}, \quad i = 1,2,3,4,5 \quad (1)$$

$$\phi(L)\Delta c_t = \mu_{s_t} + v_t, \quad v_t \sim i.i.d.N(0,1), \quad (2)$$

$$\psi_i(L)e_{it} = \varepsilon_{it}, \quad \varepsilon_{it} \sim i.i.d.N(0,\sigma_i^2), \quad (3)$$

$$\mu_{s_t} = \mu_0(1 - S_t) + \mu_1 S_t \quad S_t = \{0,1\}, \quad (4)$$

$$\Pr[S_t = 1 | S_{t-1} = 1] = p, \quad \Pr[S_t = 0 | S_{t-1} = 0] = q, \quad (5)$$

where in equation (1), Δy_{it} represent the differences of the logs of the economic variables, Δc_t is the growth rate of the common factor, hence of the composite coincident index, γ_i are the individual weights measuring the variables' sensitivity to the business cycle. Equation (2) expresses the dependence of Δc_t on whether the economy is in recession or expansion.¹⁵ Hence it incorporates the asymmetry characteristic of business cycles since, according to equations (4) and (5), the common component is assumed to be generated by a two-state Markov switching

¹⁴ Otherwise the means of the processes are overdetermined and the model is not identified.

¹⁵ As in Kim and Yoo (1995), Chauvet (1995) and the application in Chapter 5 of Kim and Nelson (1999), we allow here for the intercept term of the common component to be regime dependent. In the model of Diebold and Rudebusch (1996) it is the mean of the common component which is allowed to be regime dependent.

process. The unobservable state variable S_t switches between the recessionary 0 and the expansionary 1 regimes, where the respective transition probabilities q and p are given by equation (5). The processes of the idiosyncratic components e_{it} are described by equation (3) and the innovations v_t and ε_{it} are assumed to be independent for all t and i .

For estimation purposes, the model is cast into state space form. Parameter estimates as well as the unobserved common factor and the inferred probabilities are obtained by maximum likelihood estimation. However, due to the non-linearity of the model, the usual Kalman filter is not applicable.¹⁶ The applied estimation procedure therefore consists of a combination of Hamilton's algorithm and a non-linear version of the Kalman filter, as introduced by Kim (1994).¹⁷

We select AR(2) processes for the common factor and the idiosyncratic terms. Other higher parameterized specifications were also used, but they did not provide any additional statistically significant coefficients or a significantly different estimate of the common factor. Furthermore, a second lag of Δc_t is allowed to enter the tourism equation of the model.¹⁸ For purposes of identification, all series are demeaned by subtracting the mean difference and the variance of v_t is set equal to unity. The selected model specification can be described by the following measurement and transition equations:¹⁹

¹⁶ For an intuitive explanation for that see Kim and Nelson (1999).

¹⁷ The basic filtering and also smoothing procedures involve some approximations. Kim and Nelson (1999) argue that in cases where exact maximum likelihood estimation is not feasible, the approximation-based algorithms can be employed, since the approximations appear to be quite accurate with only marginal efficiency cost. They also introduce the Gibbs sampling approach to enable approximation-free inference in non-linear models.

¹⁸ See also Section 3.

¹⁹ The GAUSS programs written to estimate the models draw on routines that are referred to and presented on the website of Kim and Nelson (1999).

$$\begin{bmatrix} \Delta y_{1t} \\ \Delta y_{2t} \\ \Delta y_{3t} \\ \Delta y_{4t} \\ \Delta y_{5t} \end{bmatrix} = \begin{bmatrix} \gamma_1 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ \gamma_2 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ \gamma_3 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ \gamma_{40} & \gamma_{41} & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ \gamma_5 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix} \begin{bmatrix} \Delta c_t \\ \Delta c_{t-1} \\ e_{1t} \\ e_{1t-1} \\ e_{2t} \\ e_{2t-1} \\ e_{3t} \\ e_{3t-1} \\ e_{4t} \\ e_{4t-1} \\ e_{5t} \\ e_{5t-1} \end{bmatrix}, \quad (6)$$

$$\begin{bmatrix} \Delta c_t \\ \Delta c_{t-1} \\ e_{1t} \\ e_{1t-1} \\ e_{2t} \\ e_{2t-1} \\ e_{3t} \\ e_{3t-1} \\ e_{4t} \\ e_{4t-1} \\ e_{5t} \\ e_{5t-1} \end{bmatrix} = \begin{bmatrix} \mu_{s_t} \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} + \begin{bmatrix} \phi_1 & \phi_2 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & \psi_{11} & \psi_{12} & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & \psi_{21} & \psi_{22} & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & \psi_{31} & \psi_{32} & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & \psi_{41} & \psi_{42} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & \psi_{51} & \psi_{52} \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix}$$

$$* \begin{bmatrix} \Delta c_{t-1} \\ \Delta c_{t-2} \\ e_{1t-1} \\ e_{1t-2} \\ e_{2t-1} \\ e_{2t-2} \\ e_{3t-1} \\ e_{3t-2} \\ e_{4t-1} \\ e_{4t-2} \\ e_{5t-1} \\ e_{5t-2} \end{bmatrix} + \begin{bmatrix} v_t \\ 0 \\ \varepsilon_{1t} \\ 0 \\ \varepsilon_{2t} \\ 0 \\ \varepsilon_{3t} \\ 0 \\ \varepsilon_{4t} \\ 0 \\ \varepsilon_{5t} \\ 0 \end{bmatrix}, \quad (7)$$

5. Empirical results

The estimated dynamic factor model with regime switching appears to successfully provide information about cyclical movements in Greek economic activity. Furthermore, the model seems to fit the data quite well, since almost all parameters have the expected signs and are statistically significant. Moreover, the empirical evidence provides support for the two-state specification, and hence, for the asymmetry between recessions and expansions.

More specifically, the two autoregressive coefficients of the common component are negative and statistically significant. As regards the idiosyncratic components, the factor loadings picturing the sensitivity of the variables to the business cycle are in all cases positive and in all but one cases statistically significant. More specifically, industrial production has the highest weighting on the common factor with a coefficient of 0.30 and, hence, seems to respond the most to business cycle changes in economic conditions. Imports, tourism and cars are less sensitive but still responsive with coefficients around 0.1. In contrast, retail sales appear to respond the less to business cycle fluctuations with a low and not statistically significant factor loading of 0.05. Hence, among the components included, retail sales is the least cyclical one. The negative coefficients of ψ_i for all series except industrial production imply that the idiosyncratic components exhibit negative serial correlation, while the respective first coefficient for the industrial production series indicates positive idiosyncratic autocorrelation.

The empirical evidence further supports the usefulness of the Markov switching feature of the model. The adequacy of the selected model is confirmed by

the signs of the estimated Markov-switching parameters μ_0 and μ_1 . In recessions, μ_0 equaling -9.3157 is significantly below zero, while during cyclical upswings μ_1 which equals 0.0935 is positive. In line with similar findings in the literature, the probability of staying in expansion ($p=0.983$) is higher than that of staying in a recession ($q=0.337$). This agrees with the suggestion that the average duration of expansions is larger than the duration of recessions. The implied expected durations of expansion calculated as $[1/(1-p)]$ and recession given by $[1/(1-q)]$ are approximately 59 and 2 months, respectively.

The combination of the dynamic factor model with the regime switching methodology makes it possible to extract the common factor as well as the probability that the economy is in a recession. The estimated monthly coincident index of economic activity in Greece is plotted in levels in Figure 1. Unfortunately, there is no officially released business cycle chronology for Greek economic activity to compare our results to. Still, as illustrated, the constructed composite coincident index of Greek economic activity seems to perform reasonably well in picturing the business cycle for the time period under consideration. Figure 2 pictures the implied filtered and smoothed recession probabilities. As can be seen and expected, Greek business cycles have not been regular and there are no two cycles which seem to be identical. The indicated differences mainly suggest that the 1970s and 1980s were characterized by considerable up- and downswings. The same does not hold for the period after the early 1990s, which appears to be significantly more stable. The constructed coincident indicator accurately reproduces the basic downward movements of the Greek business cycle: the mid-1970s recession, which coincides with the first oil price shock and was also related to domestic political developments; a short-lived early 1980s downswing following the second oil price shock, where in this case the implied probability of

recession does not exceed 0.5; the 1987 recession and, finally, the early 1990s recession also related to the spike in the price of oil.²⁰

In order to check whether the model is correctly specified, we perform some misspecification tests on the one-step-ahead forecast errors. To confirm the validity of the model, the latter should be serially uncorrelated, hence not predictable, and normally distributed. As the Q-Ljung-Box statistics indicate, there are no signs of remaining autocorrelation. However, there are some departures from the normality hypothesis. Moreover, we check the correlation between the growth of the constructed coincident index and the growth of the included variables and also GDP growth and find positive contemporaneous correlation coefficients in all cases. The highest correlation coefficients are found for industrial production and the lowest for retail sales.

6. Conclusions

In this paper we have constructed a coincident indicator of Greek economic activity for the January 1970 to December 2007 period. Using five coincident component series and estimating a dynamic factor model with regime switching, we have obtained a common factor representing the state of the economy and the related recession probabilities. The specific methodology was selected to place emphasis on the construction of the coincident indicator on the basis of a formal econometric approach and to incorporate the two key elements of business cycles.

The provided empirical evidence supports the application of the specific methodology. Co-movement among the individual components and the regime

²⁰ If one would adopt the popular informal rule that a recession is characterized by two consecutive quarters of decline in GDP, then the recession dates indicated by the model would be all confirmed.

switching property are confirmed to be reasonable characteristics of fluctuations in Greek economic activity. Moreover, stylized facts of the Greek business cycle, such as the higher volatility of the 1970s and 1980s compared to the more stable 1990s, as well as the deepest downswings characterizing Greek economic activity during the period under investigation, are captured by the model. Hence, the proposed coincident indicator can be used for a timely evaluation of current Greek economic conditions. An interesting exercise for further research would consist in the compilation of a composite coincident indicator of Greek economic activity for the time period after the 1990s. Such a project would make it possible to take into account the important role of credit expansion and, depending on data availability restrictions, to include variables representing the employment and service sectors of the Greek economy.

Table 1 **Variables description**

Name-Mnemonic	Description	Transformation
IP	Industrial production index, monthly	Seasonally adjusted
RS	Turnover index in retail trade, monthly	Deflated, seasonally adjusted
IMP	Value of imports (arrivals), monthly	Indexed, deflated, seasonally adjusted
TRSM	Total arrivals of foreigners at the Greek borders, monthly	Indexed, seasonally adjusted
CARS	Number of issued passenger cars (private use) licenses	Indexed, seasonally adjusted
GDP	Constant prices, quarterly	Indexed, seasonally adjusted, converted to monthly

Source: General Secretariat of National Statistical Service of Greece, Eurostat.

Table 2 **ADF unit root tests and Johansen test for cointegration**

	IP	RS	Variables IMP	TRSM	CARS
ADF unit root tests					
Levels	-2.754 (0.066)	-0.160 (0.941)	0.487 (0.986)	0.217 (0.974)	-0.922 (0.781)
First differences	-16.461 (0.000)	-23.917 (0.000)	-16.631 (0.000)	-12.306 (0.000)	-15.322 (0.000)
Johansen cointegration test					
Trace test	67.335 (0.078)				
Max-eigenvalue test	25.309 (0.364)				

Notes: Augmented Dickey-Fuller test statistics are provided with the corresponding p -values in parentheses. The test critical values at the 5% and 10% levels are -2.868 and -2.570, respectively. The unrestricted cointegration rank test statistics are provided with the corresponding p -values in parentheses. The 0.05 critical values for the trace and max-eigenvalue tests are 69.819 and 33.877, respectively.

Table 3 Parameter estimates of the dynamic factor model with regime switching

Common factor						
	ϕ_1	ϕ_2	μ_0	μ_1	p	q
	-0.862 (0.086)	-0.3949 (0.0691)	-9.3157 (1.1188)	0.0935 (0.1887)	0.983 (0.0038)	0.337 (0.2634)
Idiosyncratic component						
(ΔIP_t)	γ_1		ψ_{11}	ψ_{12}	σ_1^2	
	0.3025 (0.0435)	-	0.1984 (0.1256)	-0.4388 (0.069)	0.5202 (0.0541)	
(ΔRS_t)	γ_2		ψ_{21}	ψ_{22}	σ_2^2	
	0.0469 (0.0342)	-	-0.6532 (0.0463)	-0.2728 (0.0455)	0.8279 (0.0275)	
(ΔIMP_t)	γ_3		ψ_{31}	ψ_{32}	σ_3^2	
	0.1133 (0.0343)	-	-0.4483 (0.0464)	-0.1839 (0.0463)	0.8876 (0.0297)	
$(\Delta TRSM_t)$	γ_{40}	γ_{41}	ψ_{41}	ψ_{42}	σ_4^2	
	0.0822 (0.0387)	0.1638 (0.0369)	-0.5276 (0.0489)	-0.174 (0.0483)	0.8634 (0.0303)	
$(\Delta CARS_t)$	γ_5		ψ_{51}	ψ_{52}	σ_5^2	
	0.1181 (0.0351)	-	-0.4288 (0.046)	-0.2455 (0.0461)	0.8862 (0.0297)	
Log likelihood: -807.181						

Note: Standard errors are given in parentheses.

Figure 1 The coincident indicator of Greek economic activity

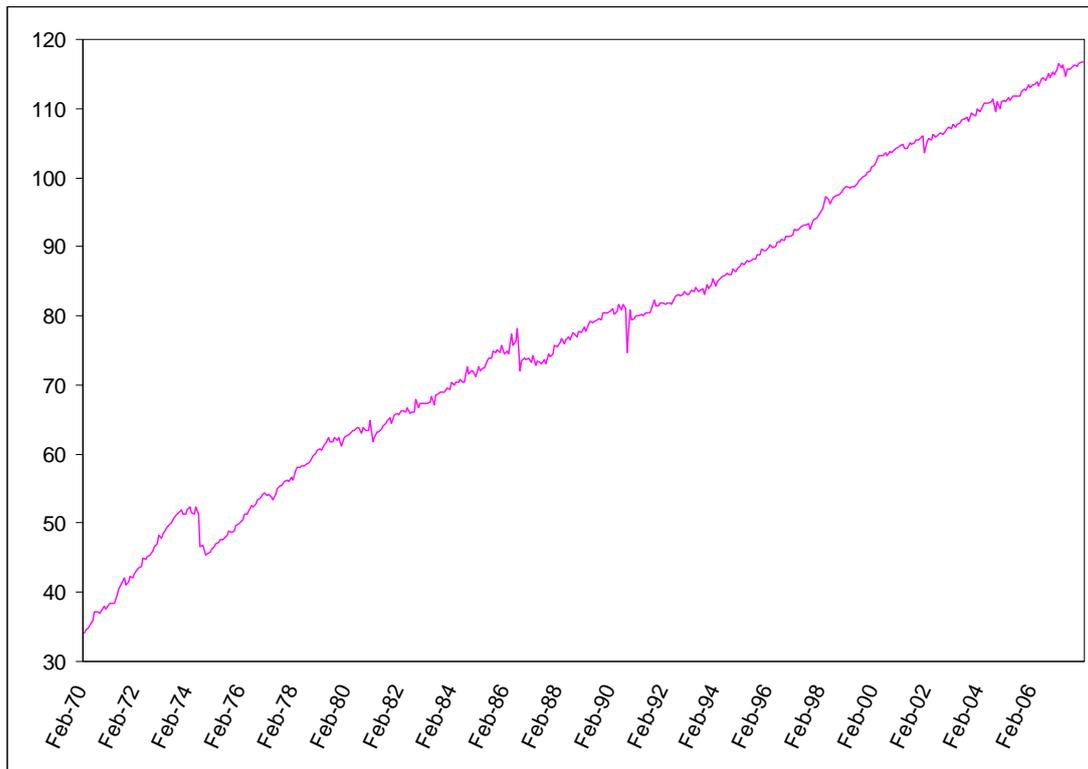
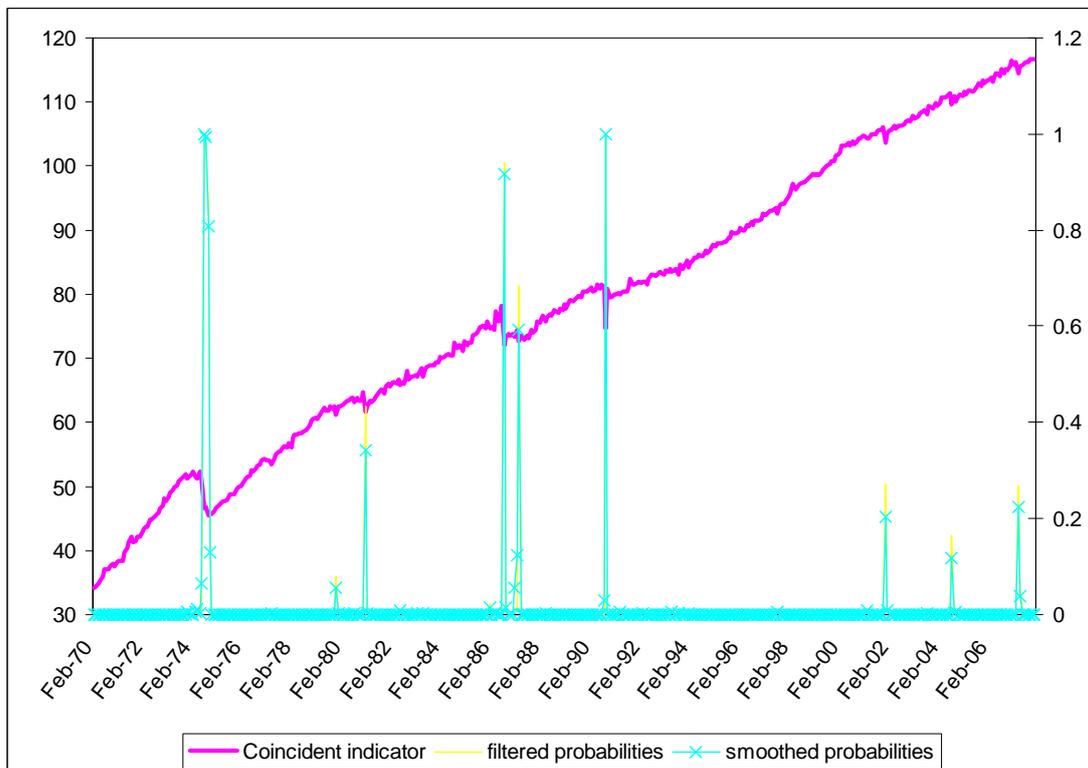


Figure 2 The coincident indicator, filtered and smoothed recession probabilities



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